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## **Title page**

**How *should* adult handgrip strength be normalized? Allometry reveals new insights and associated reference curves.**

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### **Running title**

How should handgrip strength be normalized?

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## Abstract

**Introduction:** Handgrip strength (HGS) is an important indicator of health. Because HGS is strongly associated with body size, most investigators normalize HGS for some measure of body size as a more sensitive indicator of strength within a population. We aimed to (1) identify the optimal body size dimension to remove (normalize) HGS for differences in body size among adults, and (2) generate norm-referenced centiles for HGS using the identified body size dimension.

**Methods:** Data were from the National Health and Nutrition Examination Survey (NHANES), a representative sample of the U.S. non-institutionalized civilian population. Exclusions resulted in a final sample of 8690 adults aged 20 years and older. HGS was measured using handheld dynamometry. Body size dimensions included body mass, height, and waist circumference. The most appropriate dimension(s) associated with HGS were identified using allometry. We fitted centile curves for normalized HGS using the Generalized Additive Model for Location, Scale, and Shape (GAMLSS).

**Results:** Findings suggest that neither body mass nor body mass index is appropriate to normalize HGS. Incorporating all three body size dimensions of body mass, height, and waist circumference, or the reduced sub-sets of body mass and height, or height alone, suggest that the most appropriate normalizing (body size) dimension associated with HGS should be a cross-sectional or surface area measure of an individual's body size (i.e.,  $L^2$ , where L is a linear dimension of body size). Given that height was also identified as the signally best body size dimension associated with HGS, we recommend HGS be normalized by height<sup>2</sup> (i.e.,  $HGS/HT^2$ ). Centile curves for  $HGS/HT^2$  by age group and gender were therefore provided.

**Conclusion:** Scaling HGS by height<sup>2</sup> may help normalize strength for population-based research.

**Key words**

Hand strength, body size, adult, nutrition surveys, cross-sectional studies, waist circumference.

## 1 **Introduction**

2 Muscle strength, assessed by handgrip strength (HGS) using isometric dynamometry, is  
3 considered a powerful marker of current and future health (1-5). Low adult HGS is significantly  
4 associated with an increased risk of all-cause, cardiovascular, and non-cardiovascular mortality  
5 (3, 6), stroke (3), several cancers (including colorectal, lung, and breast cancer) (6), chronic  
6 obstructive pulmonary disease (6), type 2 diabetes (7), fractures (8), cognitive declines  
7 (including dementia) (8), and functional disability (9). Low HGS is also part of decision  
8 algorithms and assessment criteria for determining sarcopenia (10), dynapenia (11), and frailty  
9 (12). HGS is easy, affordable and safe to assess (13), has moderate-to-high construct validity  
10 with total body and knee extensor strength (14), and high-to-very high test-retest reliability (15).  
11 It is for these reasons, why HGS is widely used to determine strength capacity in clinical and  
12 epidemiological settings, and for population health surveillance (16).

13  
14 However, HGS is strongly and positively associated with body size, with taller and/or heavier  
15 individuals having greater HGS. For this reason, most investigators report HGS both in absolute  
16 units (usually kg) and normalized for some measure of body size, as a more sensitive indication  
17 of strength capacity within a population where sub-groups are known to vary in body size (e.g.,  
18 gender, race). Various normalizing methods have been used to adjust HGS for differences in  
19 body size. Most investigators have normalized HGS to body mass (17-24), some have  
20 normalized to body mass index (BMI) (23-26), while few have normalized to other measures of  
21 body size (e.g., height) (22, 23). The process of normalizing variables such as HGS per body  
22 mass (an index known as a ratio standard) has come under strong criticism in the past, a point  
23 originally made by Tanner (27) and subsequently by Nevill et al. (28). Indeed, focusing on

24 scaling HGS specifically, Kiilkamp et al. (29) confirmed that HGS should not be normalized by  
25 dividing HGS by the entire body mass in both judo athletes and non-athletes. Nevertheless, such  
26 ratio standards, using HGS per body mass or HGS per BMI to normalize HGS data, have been  
27 used to develop nationally-representative norm-referenced centiles (21, 25) and criterion-  
28 referenced health-related cut-points (17, 23). This inconsistency in normalization approaches  
29 prompts the obvious research question, “How should HGS be normalized for differences in body  
30 size?”

31  
32 Hence, the purposes of the current study are twofold. Using a nationally representative sample of  
33 Americans aged 20 years and older, we aimed to (1) identify, using allometric scaling, which  
34 body size dimension is optimal to remove (adjust/normalize) HGS for differences in body size,  
35 and (2) to generate norm-referenced centile data for normalized HGS estimated using the  
36 generalized additive model for location, scale, and shape (GAMLSS) (30). We hypothesized that  
37 the most appropriate body size dimension associated with HGS was likely to be a cross-sectional  
38 area of body size such as body mass ( $M^{0.67}$ ), see for example, Kiilkamp et al. (29) when  
39 normalizing HGS and Nevill et al. (28) when normalizing maximal oxygen uptake for  
40 differences in body size.

41

## 42 **Methods**

### 43 *Participants*

44 We used data from the 2011-12 and 2013-14 cycles of the National Health and Nutrition  
45 Examination Survey (NHANES) dataset, which used a complex multistage probability design to  
46 assess the health and nutrition status of a representative sample of the U.S. non-institutionalized

47 civilian population (31). These cycles of the NHANES were selected because they included  
48 measures of HGS. Written informed consent was provided by participants and the National  
49 Center for Health Statistics Research Ethics Review Board approved NHANES protocols  
50 (Protocol #2011-17). We did not seek additional approval because the data used in this  
51 study were free from personal identifiers.

52

53 While NHANES recruited participants aged 6 years and older, we only used data on adults aged  
54 20 years and older (20-80+ years, with adults aged 80 years and over top-coded in the NHANES  
55 at 80 years of age) in this study. Of the initial 19,931 participants, 10,988 were excluded because  
56 they: (a) were younger than 20 years ( $n=8602$ ), (b) were pregnant ( $n=174$ ), (c) performed the  
57 HGS assessment seated (due to physical limitations) 386), (d) were not assessed for HGS with  
58 both hands ( $n=1539$ ), or (e) had missing data (e.g., body mass, height, waist circumference)  
59 ( $n=287$ ). In addition, following the procedures of Wang et al. (32) we excluded a further 253  
60 participants as outliers because their bilateral HGS asymmetry was  $>30\%$ . These exclusions  
61 resulted in a final sample of 8690 adults aged 20 years and older.

62

### 63 *Measures*

64 The HGS and anthropometry protocols are described in detail elsewhere (33-36). HGS was  
65 measured using Takei digital handgrip dynamometer (Model T.K.K.5401, Takei Scientific  
66 Instruments, Niigata City, Japan). Participants were randomly assigned to start the HGS test with  
67 their right or left hand, with the dynamometer adjusted for hand size by ensuring that the middle  
68 phalange of each participant's index finger was bent to  $90^\circ$  and rested flat atop of the handle. A  
69 sub-maximal effort practice trial was performed to ensure the dynamometer was properly

70 adjusted for hand size and to confirm understanding of the HGS protocol. Participants stood  
71 upright (unless they were physically limited), with their feet hip width apart, their arm extended  
72 and hanging down away from their body, and squeezed the dynamometer with maximal effort.  
73 Three trials were performed for each hand, alternating hands between trials, with 60 seconds of  
74 rest between measures on the same hand. The coefficient of variation across the three trials was  
75 8.2%, equivalent to a typical error of 2.8 kg. For this study, HGS was taken as the average of the  
76 maximum score attained for each hand.

77  
78 Standing height was measured using a fixed stadiometer with an adjustable headboard. Body  
79 mass was measured using a Mettler Toledo digital weight scale (Mettler-Toledo, Columbus, OH,  
80 USA). Waist circumference was measured at end-tidal expiration using a steel measuring tape  
81 placed directly on the skin at the level of the superior lateral border of the iliac crests.  
82 Participants self-reported their age and gender.

83  
84 *Statistical analyses*

85 To obtain nationally representative estimates, analyses were conducted using NHANES sample  
86 weights (survey, strata, and cluster weights), which account for the complex survey design  
87 (including oversampling), survey non-response, and post-stratification. To identify the most  
88 appropriate body size dimension(s) associated with HGS, we developed the following  
89 multiplicative model with allometric body size components, similar to that used to model the  
90 physical performance variables of Greek children (37), Peruvian children (38), and older adults  
91 (39).

$$92 \quad HGS = a \cdot M^{k_1} \cdot HT^{k_2} \cdot WC^{k_3} \cdot \varepsilon$$



93 where ‘a’ is the scaling constant and  $k_1, k_2$ , and  $k_3$  are scaling exponents for the body mass (M),  
94 height ( *HT*), and waist circumference ( *WC* ) respectively, and  $e$  is the multiplicative error ratio  
95 (28). Note that the multiplicative error ratio ‘s’ assumes that the error will increase in proportion  
96 to body size, a characteristic in data known as heteroscedasticity that can be controlled by taking  
97 logarithms, as described below. Age and gender were incorporated into the model by allowing ‘a’  
98 to vary for either gender and each age group (age categories 20-29 years, 30-39 years, ..., 80 years  
99 and over) to accommodate the likelihood that HGS may rise and then peak sometime during  
100 middle age and decline thereafter. The model can be linearized with a log-transformation, and  
101 multiple regression/ANCOVA can be used to estimate the body mass and height exponents for  
102 HGS having controlled for both age and gender (Eq. 2). In effect, log-transformed HGS becomes  
103 the dependent variable, with age and gender incorporated as fixed factors with  $\log(M)$   
104 and  $\log(HT)$  entered as the covariates.

$$105 \quad \log(HGS) = \log(a) + k_1 \cdot \log(M) + k_2 \cdot \log(HT) + k_3 \cdot \log(WC) + \log(\varepsilon) \quad (2)$$

106 Traditionally,  $R^2$  is used to measure goodness of fit. However, higher  $R^2$  values do not always  
107 indicate a better fit. Higher  $R^2$  can indicate overfitting and adding noise variables will also inflate  
108  $R^2$ . While  $R^2$  is useful, it is not necessarily the best method of comparing competing models. An  
109 alternative method of model comparison is to use the Akaike information criterion (AIC) that can  
110 be conceptualized as a “distance” or error between the data and a model, with lower values  
111 indicating a better model. Unlike  $R^2$ , which rewards models for being more complex (i.e., having  
112 more noise variables) with a higher value being better, AIC penalizes models for being more  
113 complex with a lower value being better. As a result, model comparison (goodness-of-fit) between  
114 the allometric models and the equivalent linear, additive models was assessed using the AIC. The  
115 difference between two AIC values was interpreted as negligible ( $<2$ ), moderate ( $>2$  and  $<6$ ),

116 strong (>6 and <10), or very strong (>10) evidence for the model with the lower AIC value being  
117 better.

118

### 119 *Methods for developing the centile curves*

120 Using a group of models called Generalized Additive Model for Location, Scale and Shape  
121 (GAMLSS) (30), we fitted centile curves for the most appropriate normalized HGS ratio (to be  
122 identified in the results section) by age and gender. Using this approach, we were able to fit  
123 different response distributions and different nonparametric smoothing functions (cubic splines, P-  
124 splines, and local polynomial regression). The response distributions fitted included the Box- Cox-  
125 t, Box-Cox Cole and Green, and Box-Cox Power Exponential. Each model included NHANES  
126 sample weights to adjust the dependent variable for oversampling and to better estimate  
127 population parameters. We selected the best fitting models using scaled AIC values (40), which  
128 ranks models according to their relative importance. The Box-Cox-t (p, o, v, x) power  
129 transformation produced the best fit for both males and females. This distribution was defined by  
130  $Y_v$  having a shifted truncated t distribution with x degrees of freedom, is a four- parameter  
131 distribution, which includes location (p) the median of the distribution, scale, sigma (a),  
132 approximately the coefficient of variation,  $nu$  (v) the controls for skewness (the transformation to  
133 symmetry), and  $tau$  (x) the kurtosis of the distribution (30).

134

135 The effects and covariates assessed using the ANCOVAs were considered significant at  $p < 0.05$ .  
136 All statistical analyses were conducted in IBM SPSS Statistics (v26, IBM, Chicago, IL, USA),  
137 except for the centile curves which were conducted in R (v4.0.2 (41)). We used the GAMLSS  
138 package to fit centile curves (30). Post-estimation diagnostics for these models included standard

139 QQ-plots, de-trended normal QQ-plots (worm plots (42)) and transformed Owen's plots, to check  
140 the age-conditional normality of the transformed data (43).

141

## 142 **Results**

143 To illustrate the strong and positive association between HGS and body size ( $r=0.73$ ,  $p<0.001$ ),  
144 the HGS of U.S. men and women were plotted against height in Figure 1. This figure provides  
145 evidence that the errors increase with height, a characteristic in data known as heteroscedasticity  
146 that can be controlled by taking logarithms as described previously in the methods section.

147 \*\*\*Insert Figure 1 here\*\*\*

148 The ANCOVA analysis of log-transformed HGS identified the main effects of gender and age as  
149 significant (age and gender; both  $p<0.001$ ) but not the age-by-gender interaction ( $p>0.05$ ). The  
150 main effects of age and gender are shown in Figure 2.

151 \*\*\*Insert Figure 2 here\*\*\*

152 The ANCOVA analysis also revealed that all three body size covariates were significant (Table  
153 1). Note that fitted body mass ( $M$ ) and height ( $HT$ ) exponents are both positive but waist  
154 circumference ( $WC$ ) is negative, confirming that greater body mass and height benefit HGS but  
155 excess waist circumference is detrimental to HGS.

156 \*\*\*Insert Table 1 here\*\*\*

157 If waist circumference were unavailable, the reduced body mass and height allometric model  
158 covariates for HGS are given in Table 2.

159 \*\*\*Insert Table 2 here\*\*\*

160 Finally, examining the log-transformed body size covariates in Table 2, the height covariate  
161  $\log(HT)$  appears to be the dominant body size dimension associated with HGS ( $t$  score is nearly

1 6 2 twice as large as that associated with body mass). For this reason, we re-ran the ANCOVA  
1 6 3 analysis, using a simplified/reduced allometric model (Eq. 2), excluding  $\log(M)$  and  $\log(WC)$ .  
1 6 4 The follow-up analysis revealed the height covariate as highly significant (Table 3), but the fitted  
1 6 5 height ( $HT$ ) exponent was very close to 2 (i.e.,  $HT^2$ ), suggesting that if we were to use height  
1 6 6 alone, HGS should be normalized by dividing HGS by  $HT^2$ .

1 6 7 \*\*\*Insert Table 3 here\*\*\*

1 6 8 Note that the simplified ANCOVA analysis of log-transformed HGS also confirmed very similar  
1 6 9 age and gender main effects as those reported Figure 2, with main effects for both age and gender  
1 7 0 significant ( $p < 0.001$ ) but not the age-by-gender interaction ( $p > 0.05$ ) (Figure 3).

1 7 1 \*\*\*Insert Figure 3 here\*\*\*

1 7 2 To assess the benefit of using allometric scaling to determine the appropriate body size  
1 7 3 dimension to normalize HGS as independent of body size, we calculated the AIC for the above  
1 7 4 log model 3 and compared it with the AIC obtained from fitting the equivalent linear, additive  
1 7 5 models using Height and Height<sup>2</sup> as covariates. The AIC for the allometric (log-linear) model 3  
1 7 6 was 57256. When we fitted the equivalent linear additive model to predict HGS (using the fixed  
1 7 7 factors gender and age group plus the gender-by-age group interaction) but allowing height or  
1 7 8 height<sup>2</sup> as the covariates, the AICs were 58333 and 58312, respectively. Clearly the AIC  
1 7 9 associated with allometric Model 3 (AIC=57256) is vastly superior to the equivalent linear,  
1 8 0 additive models AIC =58333 and 58312 (differences >1000), respectively, evidence for very  
1 8 1 strong differences.

1 8 2  
1 8 3 The centile curves for the  $HGS/HT^2$  by age are given for males and females separately in Figure  
1 8 4 4.

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\*\*\*Insert Figure 4 here\*\*\*

These curves enable the reader to estimate an individual's normalized HGS ( $HGS/HT^2$ ) using a nationally-representative sample of American adults for comparative purposes. These centile curves provide a straightforward interpretation and add a valuable level of precision. For example, in the case of an individual's  $HGS/HT^2$  slope and age, if their estimate is on the 75<sup>th</sup> centile, it means that for every 100 individuals of the same age, 75 would have a lower  $HGS/HT^2$  slope and 25 a higher  $HGS/HT^2$  slope. Point-estimate centile tables by age for males and females are also given in Table 4.

\*\*\*Insert Table 4 here\*\*\*

**DISCUSSION**

Our initial findings, obtained by fitting the multiplicative allometric model log-transformed (Eq. 2) with all three body size terms, suggest that to obtain a normalized  $HGS_n$  independent of body mass, height, and waist circumference, we need to calculate the normalized ratio

$$HGS_n = HGS / (HT^{0.968} \cdot M^{0.577} \cdot WC^{-0.619}) \tag{3}$$

see the exponents reported in Table 1. Physiologically this finding makes perfect sense. Taken together, the body mass and waist circumference terms ( $M^{0.577} \cdot WC^{-0.619}$ ) suggest a body mass divided by WC ratio, where the latter is likely to reflect a measure of adiposity providing a ratio likely to be a proxy for lean body mass. The height term will reflect an advantage that a taller individual will be able to exert on the handheld dynamometer, probably due to the mechanical advantage of having longer levers.

207 The fitted exponents in the model (Table 1) are also entirely compatible from dimensional  
 208 considerations, as anticipated by Astrand and Rodahl (44). In their chapter on body dimensions  
 209 and muscular exercise, Astrand and Rodahl (44) reported that force should scale to the  
 210 physiological dimension of  $L^2$ , where  $L$  is a linear dimension of body size. Using  $L$  as the  
 211 common linear body size dimension (e.g., body mass,  $M = L^3$ ), the HGS denominator becomes  
 212  $(HT^{0.968} \cdot M^{0.577} \cdot WC^{-0.619}) = L^{0.968} \cdot (L^3)^{0.577} \cdot L^{-0.619} = L^{0.968} \cdot L^{1.731} \cdot L^{-0.619} = L^{2.086}$ , or  
 213 approximately  $L^2$ . This is equivalent to a body surface or cross-sectional area, suggesting that  
 214 HGS is associated with, or proportional to, muscle cross-sectional area. Many muscle  
 215 physiologists might well have anticipated and approved of this dimensional interpretation.

216  
 217 Based on the reduced body mass and height allometric model (Table 2), to obtain a normalized  
 218  $HGS_n$  independent of body mass and height, we need to calculate the normalized ratio

$$219 \quad HGS_n = HGS / (HT^{1.438} \cdot M^{0.164}) \quad (4)$$

220 The fitted exponents from the reduced-model covariates can also be interpreted from the above  
 221 dimensional considerations. The body mass and height exponents are  $(HT^{1.438} \cdot M^{0.164}) =$   
 222  $L^{1.438} \cdot (L^3)^{0.164} = L^{1.438} \cdot L^{0.492} = L^{1.93}$ , again approximately  $L^2$ .

223  
 224 Finally, using the simplified/reduced allometric model (Eq. 2), incorporating only log-  
 225 transformed height  $\log(HT)$  (excluding  $\log(M)$  and  $\log(WC)$ ), the fitted height ( $HT$ )  
 226 exponent was 1.752 (see the parameters in Table 3), again close to 2, suggesting that if we were  
 227 to use height alone, HGS should be normalized to  $HT^{1.752}$  as follows,

$$228 \quad HGS_n = HGS / (HT^{1.752}) \quad (5)$$

229 a finding that is remarkably similar to the result reported by Neto et al. (22) who recommended  
230 that HGS of older adults should be normalized using absolute HGS divided by height<sup>184</sup>.

231  
232 These results, using allometric models, suggest that the most appropriate body size components  
233 that will optimally remove the effect of body size when normalize HGS, should include all three  
234 terms body mass, height, and waist circumference, as given by Eq. 3. These results also suggest  
235 that investigators who normalize HGS using either body mass (17-24) and/or BMI (23-26) are  
236 probably using inappropriate normalizing body size terms. Clearly, if body mass and height are  
237 to be used, they should be combined by multiplying the  $M$  and terms together ( $^{38}M^{0164}$ ), see Eq.  
238 4, not dividing body mass ( $M$ ) by height ( $HT^2$ ) as is the case when using BMI ( $\text{kg.nT}^2$ ) to  
239 normalize HGS. Furthermore, if only one body size component were to be used to normalize  
240 HGS, height ( $HT^{1752}$ ) is considerably more successful than body mass ( $M$ ) at removing the  
241 body size/dimensional effect when normalizing HGS.

242  
243 We recognize that these fitted exponents adopted in the normalizing equations 3, 4, and 5 above  
244 are all “sample specific”. That is, they are likely to work well for American adults, and even  
245 though they are both physiologically and dimensionally sound, they are unlikely to be equally  
246 successful with, and generalizable to, other populations. This was illustrated perfectly when we  
247 compare the fitted denominator exponent  $HT^{1752}$  reported in Eq. 5, with the same model adopted  
248 by Neto et al. (22) for older Brazilian adults, given as  $HT^{184}$ .

249  
250 However, when normalizing HGS, we need a *simple* methodology that is likely to be  
251 “generalizable” to all populations. The one consistent and robust finding from the above

252 allometric models, was that the normalizing (body size) dimension associated with HGS was  
253 given by  $L^2$  (a cross-sectional or surface area). Furthermore, given that we were able to confirm  
254 that height ( $HT^2$ ) was the single best body size dimension associated with HGS, we recommend,  
255 in response to the question posed in the title, that HGS should be normalized by dividing HGS by  
256 height ( $HT^2$ ).

257  
258 Our findings have several implications. First, several studies investigating the associations  
259 between HGS and health have used scaling approaches that are not optimal. There is a need for  
260 future research to determine if using  $HT^2$  to normalize HGS impacts these associations. Second,  
261 to improve comparability throughout the literature, we also recommend reporting raw HGS  
262 values (i.e., in the measured units) in addition to normalized HGS values, when possible. Last,  
263 we recommend using a quintile framework to facilitate the interpretation of these HGS centiles,  
264 similar to previous studies (e.g., 45). For instance, adults below the 20<sup>th</sup> centile can be considered  
265 as having ‘very low’ HGS; between the 20<sup>th</sup> and 40<sup>th</sup> centiles as having ‘low’ HGS; between the  
266 40<sup>th</sup> and 60<sup>th</sup> centiles as having ‘moderate’ HGS; between the 60<sup>th</sup> and 80<sup>th</sup> centiles as having  
267 ‘high’ HGS; and above the 80<sup>th</sup> centile as having ‘very high’ HGS.

268  
269 **Conclusions**  
270 HGS is considered an important indicator of health. However, because HGS is strongly  
271 associated with body size, most investigators report HGS normalized for some measure of body  
272 size as a more sensitive indication of strength capacity of individual within a population. Some  
273 investigators choose to normalize HGS per unit of body mass (kg) whilst others normalize per  
274 unit BMI ( $\text{kg}\cdot\text{m}^{-2}$ ). The current study suggests that neither body mass nor BMI are appropriate to



275 normalize HGS. Incorporating all three body size dimensions of body mass, height and waist  
276 circumference, or the reduced sub-sets of body mass and height, or height alone, suggests that the  
277 most appropriate normalizing (body size) dimension associated with HGS should be a cross-  
278 sectional or surface area measure of body size (i.e.,  $L^2$ , where  $L$  is a linear dimension of body size).  
279 Given that height was also identified as the signally best body size dimension associated with  
280 HGS, we recommend HGS be normalized by dividing HGS by height<sup>2</sup> ( $HGS/HT^2$ ). For this  
281 reason, the centile curves for the  $HGS/HT^2$  by age (20-29 years, 30-39 years, ..., 80 years and  
282 over) are given separately for males and females in the study. Future research should confirm  
283 these results in other countries, preferably using nationally-representative data.

284

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287 The authors declare no conflicts of interest or external funding or sponsorship related to this  
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289 Sports Medicine. The results of this study are presented clearly, honestly, and without  
290 fabrication, falsification, or inappropriate data manipulation.

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## Figure captions

**Figure 1.** The association between handgrip strength (kg; average of the combined maximum score attained for each hand) and height (m) by gender.

**Figure 2.** The means ( $\pm$ SE) of log-transformed handgrip strength adjusted for  $\log(M)$ ,  $\log(HT)$  and  $\log(WC)$  by age group and gender.

**Figure 3.** The means ( $\pm$ SE) of log-transformed handgrip strength adjusted for  $\log(HT)$  alone by age group and gender.

**Figure 4.** The centile curves for normalized handgrip strength (handgrip strength in kilograms divided by height in meters squared) by age and gender.

**Table 1.** The fitted parameters of the ANCOVA analysis for all three body size covariates.  
**Parameter Estimates<sup>8</sup>**

Dependent Variable: Log(HGS)

<b>Parameter</b>	<b>B</b>	<b>Std. Error</b>	<b>t</b>	<b>Sig.</b>	<b>95% Confidence Interval</b>	
					<b>Lower Bound</b>	<b>Upper Bound</b>
Intercept	3.290	0.080	41.028	<0.001	3.132	3.447
Log (M)	0.577	0.024	24.341	<0.001	0.531	0.624
Log ( <i>HT</i> )	0.968	0.049	19.672	<0.001	0.871	1.064
Log(VPC)	-0.619	0.034	-18.474	<0.001	-0.685	-0.553
Female	-0.372	0.020	-18.994	<0.001	-0.411	-0.334



**Table 2.** The fitted parameters of the ANCOVA analysis adopting the log-transformed body mass and height body size covariates.

**Parameter Estimates<sup>8</sup>**

Dependent Variable: Log(HGS)

Parameter	B	Std. Error	t	Sig.	95% Confidence Interval	
					Lower Bound	Upper Bound
Intercept	1.972	0.037	52.828	<0.001	1.899	2.045
Log(M)	0.164	0.008	20.565	<0.001	0.148	0.179
Log ( <i>HT</i> )	1.438	0.043	33.521	<0.001	1.354	1.522
Female	-0.355	0.020	-17.800	<0.001	-0.394	-0.316

**Table 3.** The fitted parameters of the ANCOVA analysis adopting the log-transformed height body size covariate alone.

**Parameter Estimates<sup>8</sup>**

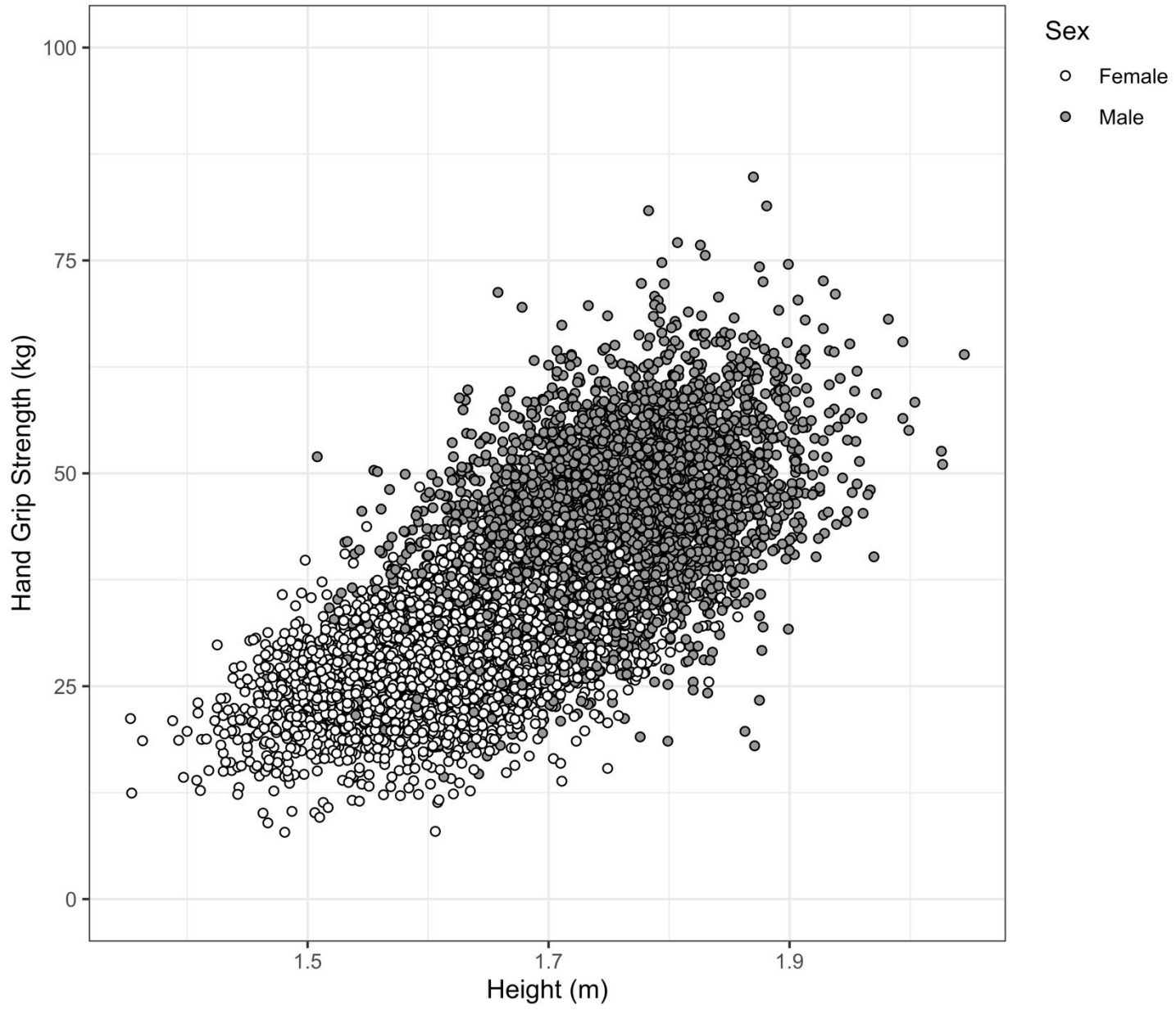
Dependent Variable: log(HGS)

Parameter	B	Std. Error	t	Sig.	95% Confidence Interval	
					Lower Bound	Upper Bound
Intercept	2.516	0.027	93.261	<0.001	2.463	2.569
Log ( <i>HT</i> )	1.752	0.041	42.678	<0.001	1.672	1.833
Female	-0.363	0.020	-17.781	<0.001	-0.403	-0.323

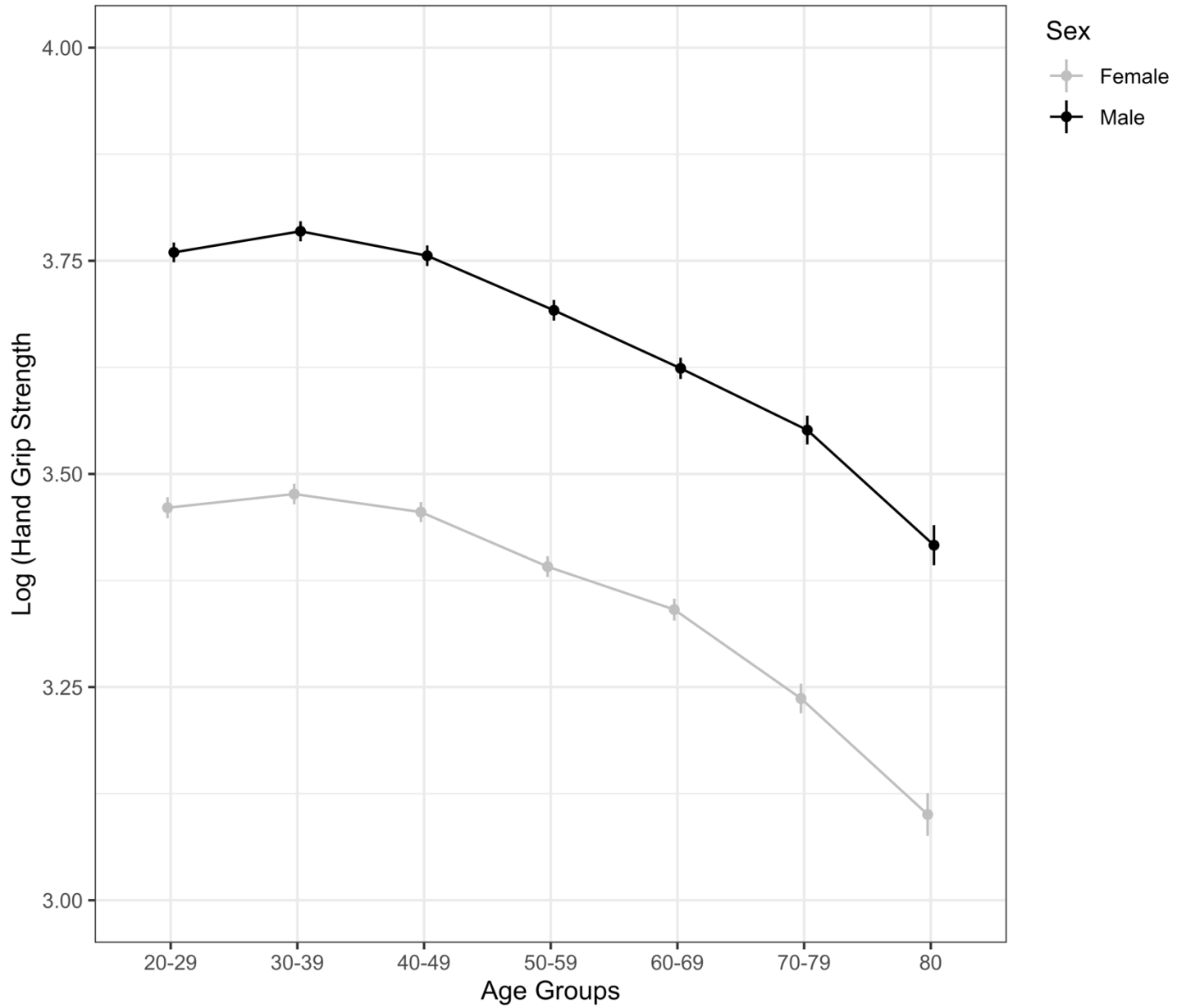
**Table 4.** Normalized handgrip strength (handgrip strength in kilograms divided by height in meters squared) centiles by age group and gender for a national sample of Americans aged 20 years and older.

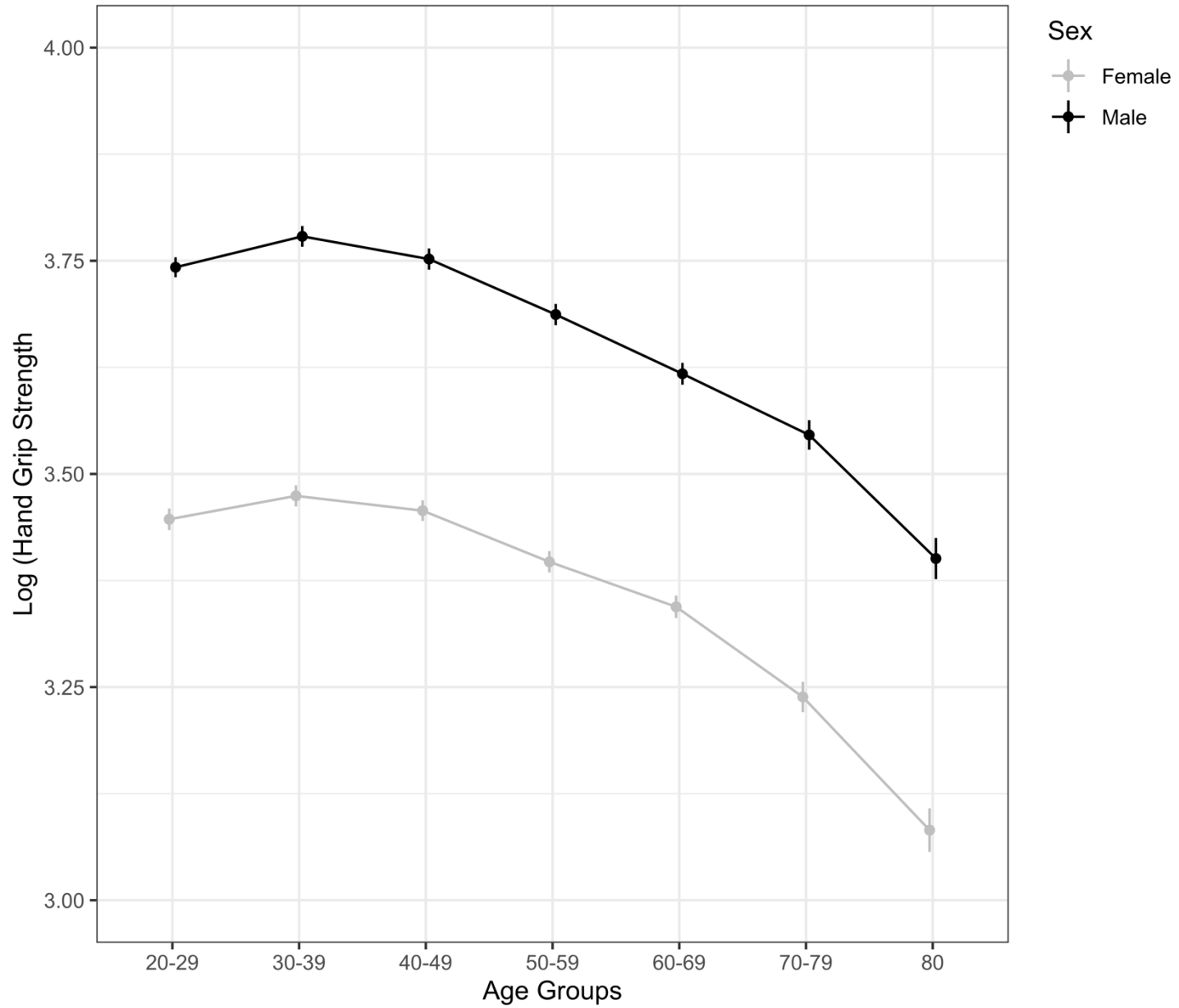
Age	C1	C3	C5	C10	C20	C30	C50	C60	C70	C80	C90	C95	C97	C99
<i>Males</i>														
20	8.4	9.5	10.0	10.9	12.0	12.7	14.1	14.7	15.5	16.4	18.3	19.0	19.8	21.6
25	8.9	10.2	10.8	11.7	12.9	13.7	15.0	15.7	16.4	17.3	18.5	19.6	20.4	21.9
30	9.2	10.6	11.2	12.2	13.4	14.2	15.5	16.1	16.8	17.6	18.6	19.7	20.4	21.8
35	9.4	10.8	11.5	12.5	13.6	14.4	15.6	16.2	16.9	17.6	18.6	19.7	20.4	21.7
40	9.5	10.8	11.5	12.4	13.5	14.2	15.4	16.0	16.6	17.3	18.4	19.4	20.0	21.5
45	9.3	10.6	11.2	12.1	13.2	13.9	15.0	15.5	16.1	16.8	17.9	18.8	19.5	20.9
50	8.9	10.2	10.8	11.7	12.7	13.4	14.5	15.0	15.5	16.2	17.4	18.1	18.7	20.0
55	8.7	10.0	10.6	11.5	12.5	13.2	14.3	14.8	15.4	16.0	16.9	17.8	18.4	19.6
60	8.0	9.4	10.1	11.0	12.0	12.7	13.8	14.3	14.8	15.5	16.5	17.3	17.8	19.0
65	6.7	8.5	9.3	10.4	11.5	12.3	13.4	13.9	14.4	15.1	16.0	17.0	17.7	19.1
70	5.7	7.7	8.6	9.8	11.0	11.7	12.9	13.4	14.0	14.7	15.3	16.8	17.6	19.4
75	4.8	6.9	7.8	9.0	10.2	11.0	12.1	12.6	13.1	13.8	14.3	16.1	16.9	19.1
80	3.8	5.8	6.8	8.0	9.1	9.8	10.9	11.3	11.9	12.6	13.3	14.8	15.7	18.1
<i>Females</i>														
20	7.4	8.0	8.4	8.9	9.5	10.0	10.8	11.2	11.7	12.2	13.0	13.7	14.2	15.1
20	7.4	8.0	8.4	8.9	9.5	10.0	10.8	11.2	11.7	12.2	13.0	13.7	14.2	15.1
25	7.7	8.3	8.6	9.2	9.8	10.3	11.2	11.6	12.0	12.6	13.4	14.1	14.5	15.4
30	7.8	8.5	8.8	9.4	10.1	10.6	11.4	11.9	12.3	12.9	13.7	14.4	14.8	15.8
35	7.7	8.5	8.9	9.5	10.2	10.7	11.5	12.0	12.4	13.0	13.8	14.5	14.9	15.9
40	7.4	8.3	8.7	9.4	10.1	10.6	11.5	11.9	12.4	12.9	13.7	14.4	14.9	15.9
45	7.0	8.0	8.5	9.2	10.0	10.5	11.4	11.8	12.2	12.8	13.6	14.3	14.8	15.9
50	6.5	7.7	8.2	8.9	9.7	10.3	11.1	11.5	12.0	12.5	13.3	14.1	14.6	15.8
55	6.2	7.4	7.9	8.7	9.5	10.0	10.9	11.3	11.7	12.2	13.0	13.7	14.3	15.4
60	6.0	7.2	7.7	8.4	9.2	9.7	10.6	11.0	11.4	11.9	12.6	13.3	13.8	14.9
65	5.8	6.9	7.4	8.1	8.8	9.4	10.2	10.6	11.0	11.5	12.2	12.9	13.4	14.5
70	5.4	6.4	6.9	7.6	8.3	8.8	9.6	10.0	10.4	10.9	11.6	12.3	12.8	13.9
75	5.0	5.9	6.3	6.9	7.6	8.1	8.9	9.2	9.6	10.1	10.9	11.6	12.2	13.3
80	4.5	5.3	5.7	6.2	6.9	7.3	8.0	8.4	8.8	9.3	10.1	10.9	11.5	12.8







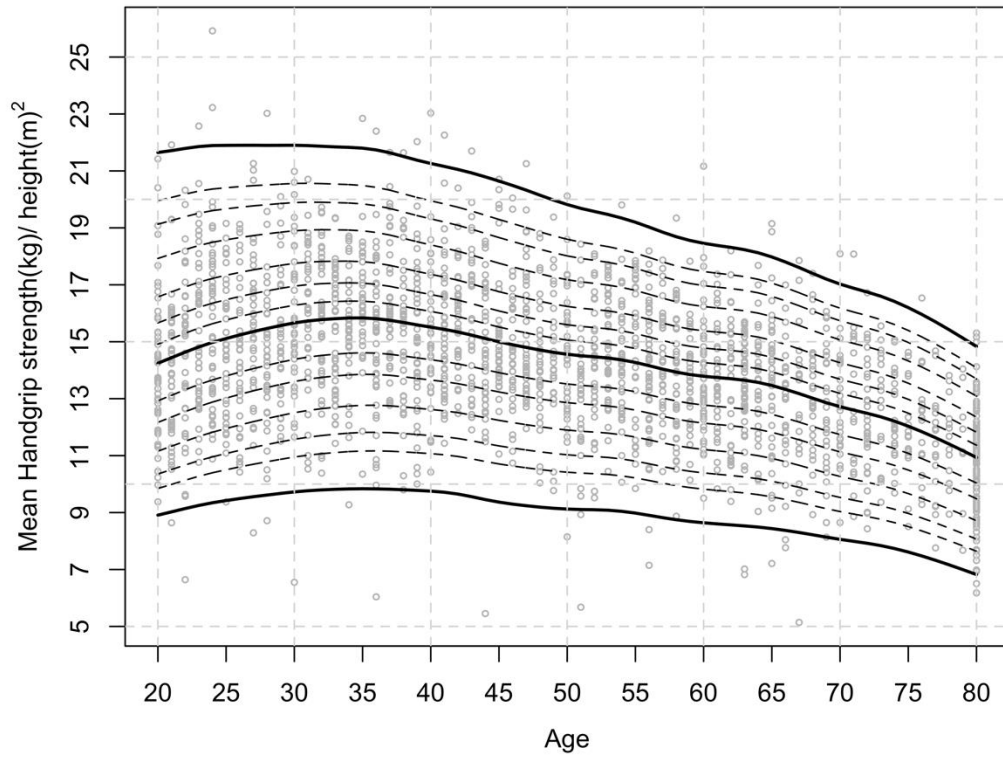








### Males



### Females

